# STAT 302: Linear Regression 

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April 15, 2020

## Transformations of Variables

Linear models are much more comprehensive than you might guess. Suppose you thought the $X, Y$ relationship was quadratic, i.e.

$$
Y_{i}=\beta_{0}+\beta_{1} X_{i}+\beta_{2} X_{i}^{2}+\epsilon_{i}
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Then just define a new covariate, $X^{2}$, by considering the squares of the $X_{i}$ !

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Then just define a new covariate, $X^{2}$, by considering the squares of the $X_{i}$ ! Interpretation: an increase of SAT from $x_{0}$ to $x_{1}$ points is associated with an expected increase in GPA of

$$
\left[\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{1}^{2}\right]-\left[\beta_{0}+\beta_{1} x_{0}+\beta_{2} x_{0}^{2}\right]=\beta_{1}\left[x_{1}-x_{0}\right]+\beta_{2}\left(x_{1}^{2}-x_{0}^{2}\right)
$$

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This cannot be interpreted on the $Y$ scale, however, because

$$
\begin{aligned}
Y & =\exp \left(\beta_{0}+\beta_{1} X+\epsilon\right) \\
& =\exp \left(\beta_{0}\right) \exp \left(\beta_{1}\right) X \exp (\epsilon)
\end{aligned}
$$

is not a linear model: it is not additive, but multiplicative!
Such models are common in, e.g., finance, because investments can be expected to grow exponentially in the long run, but noisily. When else?

## Transformations of Variables

If we suspect that the linear relationship between $X$ and $Y$ differs based on a second covariate $Z$, we can fit an interaction term:

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Y=\beta_{0}+\beta_{X} X+\beta_{Z} Z+\beta_{X Z} X Z+\epsilon
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If $Y$ is College GPA, $X$ is HS GPA, and $Z$ is sex $(0==$ male, $1==$ female), then the linear relationship between HS and College GPA can differ by sex:

$$
\begin{aligned}
& \text { (men): } Y=\beta_{0}+\beta_{X} X+\epsilon \\
& \text { (women): } Y=\left(\beta_{0}+\beta_{Z}\right)+\left(\beta_{X}+\beta_{X Z}\right) X+\epsilon
\end{aligned}
$$

## Inference

How certain are we about $\hat{Y}_{i}$ for any given $i$ ? How certain are we that a change of a unit in $X$ leads to a change of $\hat{\beta}$ in $Y$ on average? There are 2 facts we need in the univariate case:

1. Regression coefficients obey a CLT, just like the sample mean (regression with just $\beta_{0}$ and no covariates is computing the sample mean). There are assumptions...
2. $\operatorname{sd}(\hat{\beta} x)=x \operatorname{sd}(\hat{\beta})$

So if the S.E. estimate for $\hat{\beta}$ is $\hat{\sigma}$, then

1. a $95 \%$ confidence interval for $\hat{\beta}$ is $[\hat{\beta}-1.96 \hat{\sigma}, \hat{\beta}+1.96 \hat{\sigma}]$
2. a $95 \%$ confidence interval for the expected change in $Y$ associated with a change in $X$ of $a$ units is [ $a \hat{\beta}-1.96 a \hat{\sigma}, a \hat{\beta}+1.96 a \hat{\sigma}$ ]

## Multivariate inference

To get inference for $\hat{Y}_{i}$ or complex quantities in multivariate models, we need one more concept:

## Definition

Variance-covariance matrix Let $\hat{\beta}$ be a random vector (a vector of random variables). Then

$$
\operatorname{Var}(\hat{\beta})=\hat{\Sigma}=\left[\begin{array}{cccc}
\hat{\operatorname{Var}\left(\hat{\beta}_{0}\right)} & \operatorname{Cov}\left(\hat{\beta}_{0}, \hat{\beta}_{1}\right) & \cdots & \operatorname{Cov}\left(\hat{\beta}_{0}, \hat{\beta}_{d}\right) \\
\operatorname{Cov}\left(\hat{\beta}_{1}, \hat{\beta}_{0}\right) & \operatorname{Var}\left(\hat{\beta}_{1}\right) & \cdots & \vdots \\
\vdots & & \ddots & \\
\operatorname{Cov}\left(\hat{\beta}_{d}, \hat{\beta}_{0}\right) & \cdots & & \operatorname{Var}\left(\hat{\beta}_{d}\right)
\end{array}\right]
$$

All you need to know is: if $a \in \mathbb{R}^{d+1}$, then $\hat{\operatorname{Var}}\left(a^{T} \hat{\beta}\right)=a^{T} \hat{\Sigma} a$. And if lmod is a linear model object, then vcov(lmod) gives you $\hat{\Sigma}$.

